Bayesian Statistical Modeling of Spatial Error Correlations in Atmospheric Tracer Inverse Analysis

> Prasad Kasibhatla Chiranjit Mukherjee and Mike West Duke University

## 5<sup>th</sup> International GEOS-Chem Meeting May 3, 2011

Funding: NASA ACMAP and Carbon Cycle Science Program

## **Bayesian Atmospheric Tracer Inversion**

# **Bayesian Inverse Analysis**

Update prior knowledge of fluxes  $[p(\mathbf{x}]$  based on measurements  $[\mathbf{y}]$  to estimate posterior pdf  $[p(\mathbf{x}|\mathbf{y})]$  – we are usually interested in moments (e.g. mean, variance-covariance) of this posterior distribution.

# Linear problem (e.g. CO with prescribed OH)

 $y = Kx + \varepsilon$ 

where,

**K** is the Jacobian matrix derived from a CTM ( $K_{ij}$  is the contribution of a unit source at location/time = j to measurement at location/time = i).

# Gaussian assumptions

 $\epsilon \sim N [0, S_{\epsilon}]; \mathbf{x} \sim N[\mathbf{x}_{prior}, S_{prior}],$  with known  $S_{\epsilon}, \mathbf{x}_{prior}, S_{prior} \rightarrow$  the posterior pdf of interest is also Gaussian:

$$\begin{aligned} \mathbf{x} | \mathbf{y} &\sim \mathsf{N}[\mathbf{x}_{\text{post}}, \, \mathbf{S}_{\text{post}}], \\ \mathbf{x}_{\text{post}} &= \mathbf{x}_{\text{prior}} + (\mathbf{K}^{\mathsf{T}} \mathbf{S}_{\varepsilon}^{-1} \mathbf{K} + \mathbf{S}_{\text{prior}}^{-1})^{-1} \mathbf{K}^{\mathsf{T}} \mathbf{S}_{\varepsilon}^{-1} (\mathbf{y} - \mathbf{K} \mathbf{x}_{\text{prior}}) \\ \mathbf{S}_{\text{post}} &= (\mathbf{K}^{\mathsf{T}} \mathbf{S}_{\varepsilon}^{-1} \mathbf{K} + \mathbf{S}_{\text{prior}}^{-1})^{-1} \end{aligned}$$

## **Incorporating Spatial Observation Error Correlation Structure**

## **Covariance Matrix with Unknown Parameters**

We can consider  $\mathbf{S}_{\varepsilon}$  to be a function of unknown *structural parameters*  $\boldsymbol{\theta}$  as  $\varepsilon | \boldsymbol{\theta} \sim N[\mathbf{0}, \mathbf{S}_{\varepsilon}(\boldsymbol{\theta})] \rightarrow$  The estimation problem then involves estimating **x** as well as  $\boldsymbol{\theta}$ .

#### **Inverse Analysis**

 $\boldsymbol{\theta}$  can first be estimated using a maximum-likelihood approach, and  $p(\mathbf{x}|\mathbf{y})$  can be approximated as  $p(\mathbf{x}|\mathbf{y}, \boldsymbol{\theta}_{ML})$ . Alternatively,  $\boldsymbol{\theta}$  also be treated as a random variable with a prescribed prior distribution [ $p(\boldsymbol{\theta})$ ] and a fully Bayesian approach can be used to estimate the moments of the joint pdf  $p(\mathbf{x}, \boldsymbol{\theta}|\mathbf{y})$  based on a large sample generated from the joint pdf using MCMC.

#### **Computational Burden**

In either case, an iterative approach involving the calculation of  $\mathbf{S}_{\varepsilon}^{-1}$  at each iteration is needed  $\rightarrow$  this is a computationally intensive step when the dimension of  $\mathbf{S}_{\varepsilon}$  is large.

# **Conditional Autoregressive (CAR) Spatial Models**

# CAR modeling

An alternative approach that involves conditional error modeling  $\rightarrow S_{\epsilon}^{-1}$  is sparse and calculations involving  $S_{\epsilon}^{-1}$  are not computationally intensive.

# **Spatial CAR Model**

 $(\varepsilon_i | \varepsilon_j, j \neq i) \sim N[\Sigma (\rho w_{ij} / w_{i+}) \varepsilon_j, \tau_c^2 / w_{i+}],$ where,

 $w_{ij}$  are elements of a proximity matrix  $\boldsymbol{W},$  with  $w_{i+}$  =  $\boldsymbol{\Sigma}_{j}$   $w_{ij},$  and

 $\rho$  and  $\tau_{c}$  are unknown parameters.

Under this specification, it can be shown that the joint distribution  $p(\epsilon)$  is  $\epsilon \sim N [0, U^{-1}]$ , where,

 $\mathbf{U} = \tau_c^2 (\mathbf{D}_w - \rho \mathbf{W})$ , with  $\mathbf{D}_w = \text{diag}(w_{1+}, w_{2+}, ...)$ 

 $\rightarrow$ U is sparse and calculations involving U are not computationally intensive

# **Statistical Computation**

Posterior inferences from MCMC-generated large-sample of p(  $\mathbf{x}$ ,  $\rho$ ,  $\tau \mid \mathbf{y}$ )

# **Incorporating Non-Normal Priors**

## Positive fluxes

We are often interested in inverse estimates of emissions, which are positive by definition  $\rightarrow$  e.g. fluxes of CO, CO<sub>2</sub>, etc. from vegetation fires, fossil-fuel combustion.

## **Specification of Prior Distribution**

An alternative approach is to use a truncated normal distribution for the prior  $\rightarrow x_i \sim N(x_{a,i}, S_{a,i}) \ l(x_i > t_i)$ , where l(.) is the indicator function [l(.) is 1 when (.) is true and 0 when (.) is false].

Here, we choose  $t_i = x_{prior,i}/4$ , and choose  $x_{a,i}$  and  $S_{a,i}$ , so that the mean of  $x_i$  is equal to  $x_{prior,i}$  and the variance of  $x_i$  is equal to  $S_{prior,i}$ .

#### Computation

Sampling from the truncated normal distribution is straightforward to implement in the MCMC algorithm.

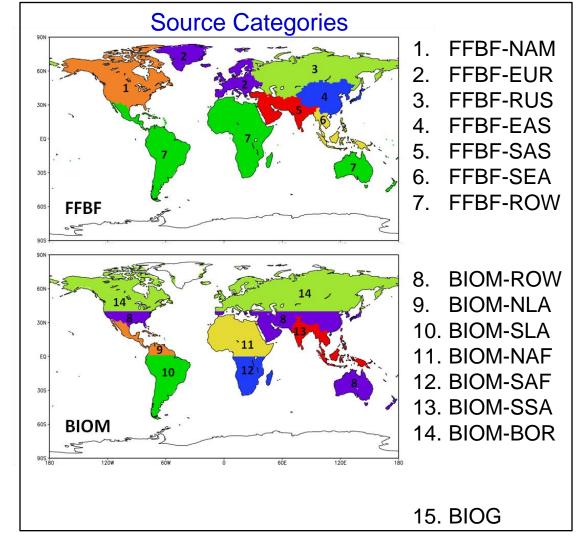
# **Synthetic Data Test**

# Synthetic Data Generation

- → Generate x<sub>true</sub> for 15 source categories by sampling from truncated normal prior
- $\rightarrow$  Generate  $\epsilon$  by sampling from

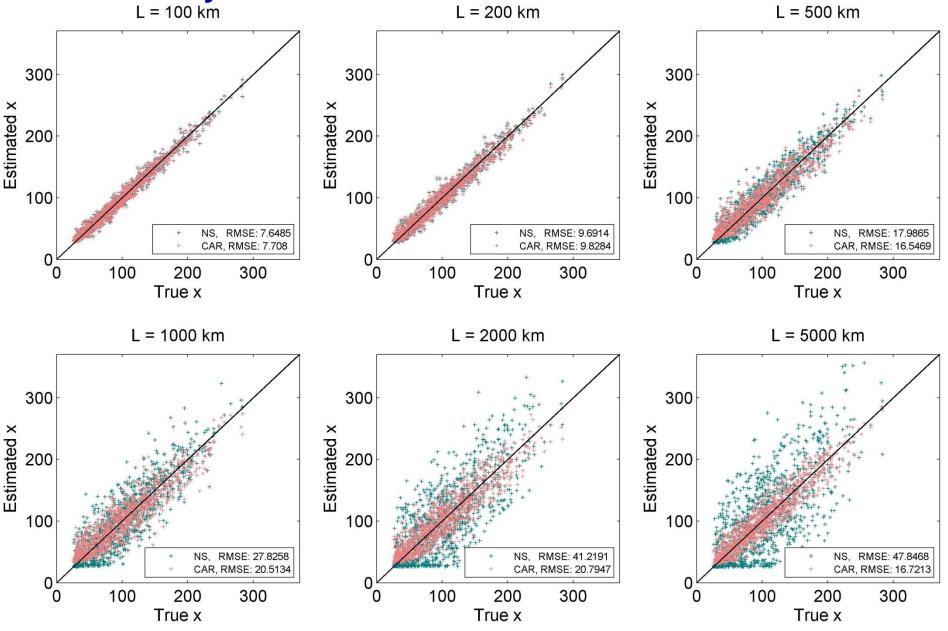
$$\begin{split} & \epsilon_{syn} \sim N \; (\textbf{0}, \, \textbf{S}_{\epsilon}) \\ & \text{with } S_{\epsilon} \left( i, j \right) = \sigma^2 \; exp(-d_{ij}/L), \\ & \text{and } \sigma = \; 20\% \; \text{of annual,} \\ & \text{global average prior model CO} \end{split}$$

- → Generate  $y_{syn} = Kx_{true} + ε_{syn}$
- → Repeat for 6 different values of L (100, 200, 500, 1000, 2000, and 5000 km)
- → Repeat for 1000 synthetic datasets for each value of L
- →K from Arellano et al. (2004) for April – December 2000



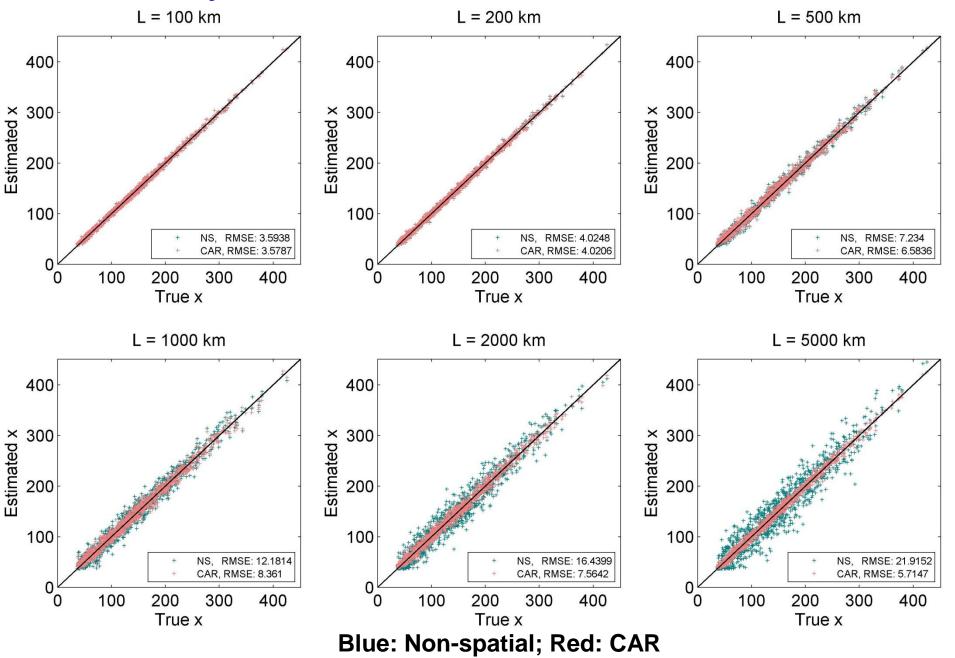
## Compare inversion results from non-spatial and CAR models

#### **Synthetic Inversion Results: FFBF-NAM**

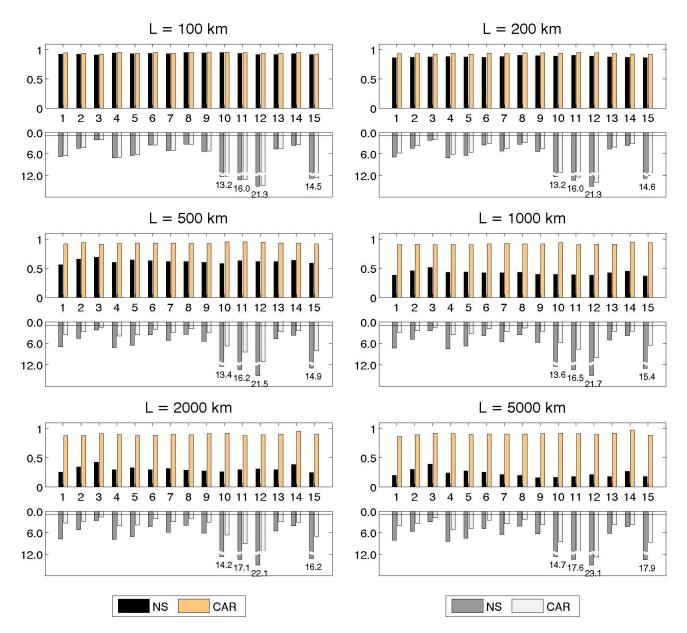


Blue: Non-spatial; Red: CAR

#### **Synthetic Inversion Results: BIOM-SAF**

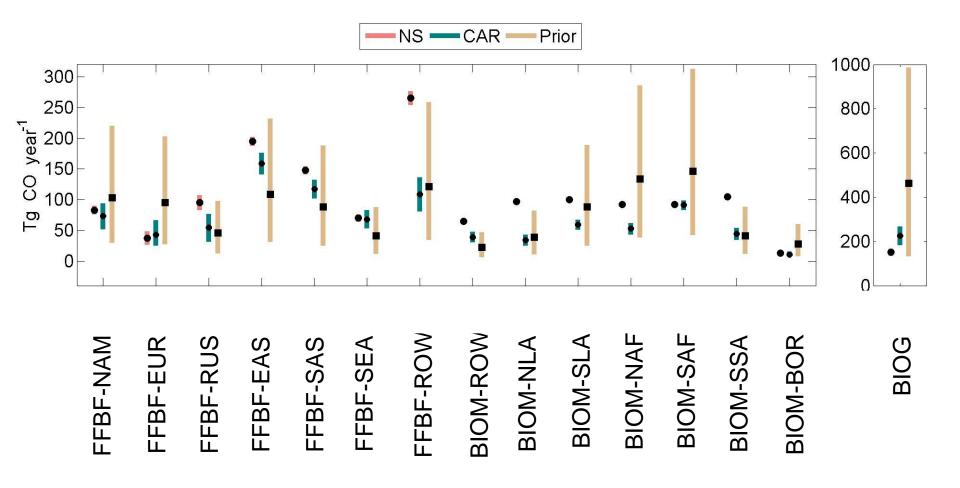


# Synthetic Inversion Results Success Rates (top) and Learning Ratios (bottom)



#### **Real MOPITT Data Inversion Results**

#### Mean (symbols) and 95% credible intervals (lines)



# Summary

- → Not accounting for spatial observation error correlations can lead to biased posterior source estimates.
- → CAR spatial modeling offers a flexible and computationally tractable approach to account for spatial error correlation structure in Bayesian atmospheric inverse modeling.

→ Future work	m	FLOPs	
Extend to multiple correlation length scales (e.g lat-dep. L)	10	70404728	
Extend to consider time correlations	100	704064128	
Integrate into an going EnKE work for grid apole inversions	1000	7040658128	
Integrate into on-going EnKF work for grid-scale inversions	10000	70406598128	